Chapter 6

Health-economic evaluation alongside a pragmatic randomised controlled trial of a web-based intervention

This chapter is published as:


Abstract

Background
Psychological interventions for the prevention of depression might be a cost-effective way to reduce the burden associated with depressive disorders.

Objective
To evaluate the cost-effectiveness of a web-based guided self-help intervention to prevent major depressive disorder (MDD) in people with subthreshold depression.

Methods
A pragmatic randomised controlled trial was conducted with follow-up at 12 months. Participants were recruited from the general population via a large statutory health insurance company and an open access website. Participants were randomised to a web-based guided self-help intervention (i.e., cognitive-behavioural therapy and problem-solving therapy assisted by supervised graduate students or health care professionals) in addition to usual care or to usual care supplemented with web-based psycho-education (enhanced usual care). Depression-free years (DFYs) were assessed by blinded diagnostic raters using the telephone-administered Structured Clinical Interview for DSM-IV Axis Disorders at 6- and 12-month follow-up, covering the period to the previous assessment. Costs were self-assessed through a questionnaire. Costs measured from a societal and health care perspective were related to DFYs and quality-adjusted life years (QALYs).

Results
In total, 406 participants were enrolled in the trial. The mean treatment duration was 5.84 (SD = 4.37) weeks. On average, participants completed 4.93 of 6 sessions. Significantly more DFYs were gained in the intervention group (0.82 vs 0.70). Likewise, QALY health gains were in favour of the intervention, but only statistically significant when measured with the more sensitive SF-6D. The incremental per-participant costs were €136 (£116). Taking the health care perspective and assuming a willingness-to-pay of €20,000 (£17,000), the intervention’s likelihood of being cost-effective was 99% for gaining a DFY, 64% or 99% for gaining an EQ-5D or a SF-6D QALY.

Conclusions
Our study supports guidelines recommending web-based treatment for subthreshold depression and adds that this may not only restore health in people with subthreshold depression, but in addition reduces the risk of developing a MDD. Offering the intervention has
an acceptable likelihood of being more cost-effective than enhanced usual care and could therefore reach community members on a wider scale.
Introduction

Currently, major depressive disorder (MDD) is one of the leading causes of years lived with disability (YLD) (1) and is further associated with substantial economic costs for society (2). In high-income countries, the 12-month prevalence of MDD is estimated at 5.1% (3) with an annual incidence rate of 3% (4). This implies that close to 60% of the depressed cases are new cases, which underscores the importance of both treatment and prevention of depression to reduce its disease burden. In fact, even when assuming full coverage of evidence-based treatments, approximately only one third of the disease burden could be averted (5, 6). In this context, attention has shifted to the prevention of MDD as an additional option to foster further reductions of depression’s disease burden by reducing the development of new cases.

Meta-analyses provide evidence for the effectiveness of psychological interventions to prevent first onsets and recurrences in depression (7-9). Especially indicated prevention, targeting subthreshold symptoms of an emerging depression, appears to be particularly effective (8). People are then screened for subthreshold depression (sD) not meeting the diagnostic criteria for a full-blown depressive disorder and are subsequently offered a preventive intervention. However, research on the cost-effectiveness of depression prevention is still limited. Economic evaluations conducted alongside randomised trials and economic evaluation studies using decision analytic modelling techniques suggest that preventive interventions for depression can be very cost-effective (10-13). Research on how to deliver such interventions on a large scale to the community is scant. Using the Internet to provide community members with effective preventive interventions is currently viewed as a potentially cost-effective way of scaling up preventive interventions (14, 15). Web-based interventions are scalable. Scalability refers to the ability of the intervention shown to be effective in a research setting to be expanded under real world conditions. To reach a greater proportion of the eligible population while retaining effectiveness, only a small increase in therapeutic resources is required. Thus, the marginal cost per additional user get lower via an economies of scale effect. Economies of scale might also reduce variable costs (i.e., therapist’s support per participant) because therapists become more efficient through better organisation and learning by doing.

To the best of our knowledge, no randomised controlled trial has investigated the cost-effectiveness of a web-based guided self-help intervention to prevent the onset of diagnosed MDD. Elsewhere we reported the primary outcomes with respect to progression to MDD at 12 months (16). Here we will evaluate the cost-effectiveness and cost-utility of the web-based intervention among self-selected members from the community suffering from subthreshold depression.
Method

Study design
We conducted and reported the health-economic evaluation in agreement with the Consolidated Health Economic Evaluation Reporting Standards (CHEERS) statement (17) and the guidelines from the ISPOR (18). We conducted the economic evaluation with a 12-month time horizon from both a societal and a public health care perspective alongside a 2-armed pragmatic randomised controlled trial in Germany to establish the cost-effectiveness and cost-utility of an indicated web-based guided self-help intervention in conjunction to usual care for people with subthreshold depression as compared to enhanced usual care (i.e., web-based psycho-education in addition to treatment as usual).

Whereas the societal perspective included all costs and consequences regardless by whom they were incurred, the health care system perspective only considers direct medical costs. Full details of the trial design can be found elsewhere (19). The study was approved by the medical ethics committee of the University of Marburg (reference number AZ 2012-35K) and registered under DRKS00004709 in the German clinical trial registry.

Participants
In total, 406 participants were included in the study. Participants were recruited from March 2013 to March 2014 from the general population via a large German health insurance company and through newspaper articles, on-air media, and related websites. We chose this open recruitment strategy as it reflects the clinical practice for this type of intervention thus enhancing the trial’s ecological validity.

Inclusion criteria were: (a) age 18 and above, (b) subthreshold depression (Centre for Epidemiologic Studies Depression Scale (CES-D) ≥ 16) as having some depressive symptoms not meeting the diagnostic criteria for a full-blown DSM-IV major depressive disorder as assessed by the telephone-administered Structured Clinical Interview for DSM-IV (SCID), (c) Internet access, and (d) informed consent. Exclusion criteria included (a) meeting DSM-IV criteria for bipolar disorder or psychotic disorder, (b) having a history of a major depressive disorder in the past six months based on Kupfer’s model (20), (c) currently receiving psychotherapy for any kind of mental health problem, (d) being on a waiting list for psychotherapy, (e) having received psychotherapeutic treatment in the past six months, and (g) showing a significant suicidal risk (item 9 of the Beck Depression Inventory > 1). As we conducted a pragmatic trial, the use of antidepressant medication was allowed as part of care-as-usual. However, participants needed to be on a stable dose for at least four weeks to be able to enter the study.
Randomisation and masking
Randomisation took place at individual level and was conducted centrally by an independent statistician not otherwise involved in the study using an automated computer-generated random numbers table. Block randomisation, of size two, was used to ensure equal sample sizes across both conditions. Details about the randomisation procedure can be found elsewhere (16).

Study participants were not masked to their treatment allocation because of the nature of the intervention. SCID interviewers were, however, unaware of participants’ randomisation status. Steps taken to maintain blinding are described in detail elsewhere (16). In case of evidence for blinding breakdown, the interviewer was changed to the second outcome interview. The research staff conducting SCID interviews were not otherwise involved in the study.

Procedures
All study participants had unrestricted access to care-as-usual (CAU). CAU for subthreshold depression entails visits to the GP, but no treatment provided by mental health specialists. If depressive symptoms deteriorate, the German S3-Guideline/National Disease Management Guideline Unipolar Depression recommends psychotherapy and the prescription of antidepressant medication (21). In our pragmatic study we did not interfere in CAU. Instead, we maintained a naturalistic CAU condition to represent current clinical practice as far as possible. It should also be noted that health care use was measured in detail (see Measures), implying that we now can describe CAU in great detail (see Table 1).

Web-based guided self-help intervention
The web-based intervention, called GET.ON Mood Enhancer, is an online multimedia interactive intervention consisting of six sessions. Each session takes about 30 minutes to complete, but the amount of time spent on a session varies among users. Participants were advised to carry out at least one, preferably two lessons per week. Intervention usage was monitored by logfile analysis. On average, participants completed 4.93 of 6 sessions. The mean treatment duration was 5.84 (SD = 4.37) weeks (16). The intervention was developed by trained psychologists and therapists at the Leuphana University. Participants created their own password to access the intervention. Trial participants used the intervention free of charge.

The intervention is based on behavioural activation and problem-solving therapy. The content of the intervention was frozen during the trial. An emphasis was placed on homework assignments to integrate newly acquired skills into daily life. Therefore, participants had the option to receive a set of about 42 standardised text-messages supporting them to integrate
the learned techniques into their lives. Participants were also supported by an online-trainer, who provided written feedback after each session and monitored adherence to the intervention. In case of non-adherence, online-trainers sent up to three reminders. The total time a trainer spent per participant was approximately three hours. Trained and supervised graduate students and health care professionals provided guidance. The guidance focussed on supporting participants to work through the exercises. Online-trainers did not provide therapeutic support. Further details about the intervention can be found in the study protocol (19).

**Enhanced usual care**

Participants in the control condition got access to an online psychoeducational intervention, which was based on the German S3-Guideline/National Disease Management Guideline Unipolar Depression (21). It informed participants about evidence-based treatments of depression should symptoms deteriorate. We thus mimicked and enhanced usual care because we provided patients with information that they may not always be offered thoroughly by their GP. Similar psycho-educational interventions have been shown to be effective in reducing depressive symptoms and is therefore suggested as a first-line intervention for early manifestations of depression in the primary care setting (22). Participants could go through the materials as often as they want. However, we did not monitor its uptake and no additional support was provided to participants in the CAU condition.

**Outcome measures**

Self-report measures (i.e., EuroQol and SF-12) were collected at baseline, post-treatment (6 weeks after randomisation), 6- and 12-month follow-up using a secured online-based assessment system (AES, 256-bit encrypted). The SCID interviews at baseline, 6- and 12-month follow-up were conducted by telephone.

**Depression-free years (DFYs)**

The main outcome in the cost-effectiveness analysis was depression-free years (DFYs). DFYs were based on the number of depression-free weeks up to the onset of a major depressive episode within the 12-month follow-up period. MDD was assessed according to DSM-IV criteria as assessed by the telephone-administered SCID (23, 24) at 6- and 12-month follow-up covering the period to the previous assessment.

Time to onset of MDD was assessed as accurately as possible using the Life Chart method as developed by Lyketsos (25). In this method, life events were recalled using a calendar method where personal landmarks are used to determine the presence of depressive symptoms.
at each month during the follow-up period. During the interview the first day of a depressive episode was established. If the exact day could not be established, the closest week (month) was defined and the mid-point of that week (month) was used. The inter-rater agreement in this study was substantial (Cohen’s kappa = 0.77).

**Quality-adjusted life years**
Quality-adjusted life years (QALYs) were used as the outcome in the cost-utility analysis. QALYs were based on the EQ-5D-3L (EuroQol (26)) and SF-6D (a subset of items of the SF-12v1 (27)). The EuroQol and the SF-12 were assessed at baseline, post-treatment (6 weeks), and 6- and 12-month follow-up.

The EQ-5D-3L comprises five items covering five domains (mobility, self-care, usual activities, pain/discomfort, and anxiety/depression), each of which is rated as causing ‘no problems’, ‘some problems’, or ‘extreme problems’. Theoretically, the EQ-5D-3L generates 243 different health states. Preference-based utilities for each of these health states are available for various countries with “full health” and “death” being anchored at 1 and 0, respectively.

The SF-6D contains 6 dimensions (each with between 2 and 5 levels) and includes 6 items of the SF-12. The SF-6D generates 7,500 different health states. Utility values were derived using Brazier’s algorithm (28, 29).

QALY health gains were estimated by calculating the area under the curve (AUC) of linearly interpolated EQ-5D-3L and the SF-6D utilities to cover the whole follow-up period of 12 months. This method weighs the 12-month period by the respective utilities of each time period. A QALY gain of 1 would thus indicate full health throughout the 12-month trial period. For the main analysis we used the EQ-5D-3L QALY based on the UK tariffs (30). For the sensitivity analyses we used the SF-6D QALYs because these are known to be more sensitive to changes in mild conditions (31).

**Resource use and costing**
We used the Trimbos and iMTA questionnaire for costs associated with psychiatric illness (TiC-P) (32, 33) for collecting data on health care utilisation and productivity losses in patients with mental health conditions. We adapted the TiC-P for use in Germany and used it at baseline and at 6- and 12-month follow-ups.

We computed health care costs, the patient’s out-of-pocket costs, the costs for informal care provided by the family and friends of the patient, and costs stemming from productivity losses due to absenteeism and lesser productivity while at work (presenteeism). Costs were expressed in Euro and indexed for the year 2013 (index factor 1.04) based on the German consumer price index (34). Costs were converted to pound sterling (£) using the
purchasing power parities reported by the Organisation for Economic Cooperation and Development (35). For the reference year 2013, €1 was equated to £0.85.

**Health care costs**

We used two German guidelines for calculating healthcare costs (36, 37). A list of unit cost prices (i.e., outpatient care) was used to compute the total health care costs on a per-participant basis (37). Unit cost prices were as follows: €20.92 for a visit to the GP, €68.06 for an internal medicine consult, €46.55 for a session with a psychiatrist, and €81.44 for a session with a psychotherapist. Hospital stays were computed at €335.52 for an in-patient day in a mental hospital and €306.41 for an in-patient day in a hospital for psychosomatic medicine and psychotherapy. Costs were estimated by multiplying the units of resource use with corresponding unit cost prices.

**Medication**

The costs of prescribed medication were based on the German drug registry (Rote Liste (38)). The basis for calculating costs of prescribed medication is the pharmacy retail price taking into account a specific pharmacy and manufacture’s discount. The rates of discount vary between private and statutory health insurances (37). Therefore, we weighted the mean costs of the three largest packages with the same agent based on the daily defined dose by the statutory population share (89% of the German population are statutorily insured).

**Intervention costs**

The total costs for the web-based cognitive-behavioural intervention were estimated at €299 (£254) per participant by the provider (GET.ON Institute) including €180 for providing feedback by online-trainers. The total cost of the intervention was based on the actual market price of this intervention that has been determined by the provider. The GET.ON Institute aims to transfer scientific knowledge related to the present research into routine health care. This institute licenses the intervention under study from the Leuphana University, Lueneburg, to provide the intervention within routine preventive services of health insurance companies in Germany. We assume that the cost of IT servers and infrastructure will increase if the intervention is scaled up because more servers are needed. However, these fixed costs are subject to an economies of scale effect (i.e., marginal costs drop per additional user). The per-participant costs for the psycho-educational intervention were estimated at €10 (£9).

**Patient and family costs**

Out-of-pocket costs were directly obtained from participants. Costs for traveling were valued at €0.30 per kilometre. Opportunity cost (i.e., time spent on the intervention) was valued at €23.10 per hour. Costs of informal care were valued using a shadow price of €18.33 per hour (37).
Costs of productivity losses

We followed the human capital approach to value productivity losses (39). Productivity losses can be caused by days not worked (absenteeism) and by reduced efficiency while at work (presenteeism). Lost workdays due to absenteeism were valued at the corresponding gross average of participants’ income per day. Lost workdays due to presenteeism were computed by taking into account the number of work days for which the participant reported reduced functioning weighted by an inefficiency score for those days. Productivity losses from unpaid work (i.e., domestic tasks) were valued using a shadow price of €18.33 per hour needed to pay for domestic help (37).

Analysis

The study was powered to demonstrate a risk reduction of 10% between study conditions as statistically significant in a survival analyses with $\alpha < 0.05$ (2-tailed), a power of $(1-\beta) = 0.80$ using survival analysis and accounting for a 20% dropout (calculated using PASS 12). However, the study was not powered to statistically test differences in health economic outcomes. Therefore, we took a probabilistic decision-making approach for our health economic inferences (40). We did not discount costs and effects because the analysis was restricted to a 12-month follow-up period.

While evaluating the clinical outcomes, we reported all analyses in accordance to the CONSORT statement (41). Data were analysed on an intention-to-treat basis meaning that all participants were included in the analyses as randomised. To this end, we used Cox proportional hazard regression analyses to test differences in time to onset of MDD (in weeks) between intervention and control group. Concurrent use of antidepressants was included as covariate into the Cox proportional hazard model (post hoc). As the use of antidepressants was not a predictor of the outcome, it was excluded from the final model.

To account for missing data in cost and utility data, we used the regression imputation procedure in Stata version 13 to obtain required predicted values (42). Predictors of outcome and dropout were identified by (logistic) regression analysis. Identifying predictors of outcome helped us to obtain the most likely values of the outcome whereas identifying predictors of dropout allowed us to correct for bias that might arise by differential loss-to-follow-up. We did not impute hospitalisation costs because only three participants (0.7%) were hospitalised during the 12-month follow-up period leading to instable imputations. Therefore, we re-assessed the impact of hospitalisation costs on outcomes in a sensitivity analysis. At baseline, mean EQ-5D utility values were the same in intervention and control group (both groups: 0.74, SD = 0.15). Therefore, no baseline adjustments were made when calculating QALYs. Differences in QALYs
and DFYs between the intervention and control groups were assessed using independent samples t-tests.

In the cost-effectiveness analysis, the incremental cost-effectiveness ratio (ICER) was based on the incremental costs per unit effect (DFY or QALY) gained. The corresponding equation is ICER = (Costs\text{INT} – Costs\text{CTR}) / (Effects\text{INT} – Effects\text{CTR}), where Costs are the annual per-participant costs and Effects are the DFYs (QALYs) in intervention and control group (subscripted with INT and CTR, respectively). Sampling uncertainty in the ICER was handled using nonparametric bootstrapping by resampling patient-level data to generate 2,500 simulations of the ICER. We bootstrapped the SURE model (seemingly unrelated regression equations; sureg command in Stata) to allow for correlated residuals of the cost and effect equations. Bootstrapping was used to obtain confidence intervals for cost-effectiveness ratios based on the percentile method, since parametric techniques are inappropriate for use on skewed variables and ratios. The bootstrapped ICERS were plotted on a cost-effectiveness plane with effects along the horizontal axis and costs along the vertical axis. A cost-effectiveness acceptability curve (CEAC) was graphed to assess the probability of the intervention being cost-effective at varying willingness-to-pay (WTP) ceilings. All analyses including uncertainty analyses were performed using Stata version 13 (42).

Sensitivity analyses
We tested the robustness of the outcomes of the main analysis in sensitivity analyses. Firstly, the EQ-5D-3L might suffer from ceiling effects when measuring QALY changes in people with mild conditions. Therefore, we measured SF-6D QALYs, which have been reported to be more sensitive to QALY changes in milder conditions such as subthreshold depression (31).

Secondly, we assessed the impact of in-patient care on the ICER because in-patient care is one of the main cost drivers and was surrounded by much uncertainty since only three hospital admissions occurred in the whole sample (3/406 = 0.7%). Such outliers (driving costs in the control condition) could lead to misleading results and these costs were removed in a sensitivity analysis.

Results
Sample characteristics
In total, 406 participants were enrolled in the study (n\text{INT} = 202; n\text{CTR} = 204). At post treatment, 366 participants (90.1%) were still participating. At 6- and 12-month follow-up, 325 (80.1%) and 286 (70.4%) participants completed the questionnaires, respectively. The CONSORT flowchart (of the participants through the trial) can be found elsewhere (16). Dropout rates did not differ
between experimental and control conditions except for the 12-month follow-up. Here, dropout was higher in the intervention group $[\chi^2(1, n = 406) = 8.36, p = .004)$. Study dropout was not associated with baseline depressive symptom severity or any socio-demographic factor. Participants’ characteristics at baseline are presented in detail elsewhere (16). In brief, the average participant was female, 45 years of age with an above average level of education and employed.

**Effects**

The mean depression-free survival time within the 12-month trial period was 43 weeks (95% CI 41 - 46) in the intervention group and 37 weeks (95% CI 35 - 40) in the control group (16), corresponding to 0.82 DFYs and 0.70 DFYs, respectively. The incremental effectiveness of 0.82 - 0.70 = 0.12 DFYs was statistically significant (95% CI 0.05 - 0.18; $t(404) = -3.37, p < .001$). Between-group differences in EQ-5D QALY gains were not statistically significant (intervention group: 0.78, SD 0.14 vs control group: 0.77, SD = 0.13; $t(404) = -0.99, p = .32$), but incremental SF-6D QALY gains differed significantly between study groups (intervention group: 0.71, SD = 0.08 vs control group: 0.67, SD = 0.07; $t(404) = -4.40, p < .001$).

**Costs**

At baseline, mean total costs were €483 (£411) in the intervention group and €528 (£449) in the control group, which is only a small difference of €45 (£38), indicating that randomisation had been well balanced. Table 1 presents the 12-month accumulated per-participant costs for various cost categories by study condition. The mean health care costs were higher in the intervention group as compared to the control condition. This difference can largely be explained by differences in the costs of the intervention as compared to the control condition with web-based psycho-education (€299 vs €10). Patient and family’s out-of-pocket costs and the costs stemming from changes in productivity losses remained quite similar between study groups. Mean total costs, as seen from both the health care and societal perspective, were hence slightly higher in the intervention group as compared to the control group (societal perspective: €143 (£121); health care perspective: €136 (£116)).
Table 1. Mean annual per-participant costs (in €) by condition cumulative over the 12-month follow-up period (based on intention-to-treat sample, N = 406)

<table>
<thead>
<tr>
<th></th>
<th>Intervention group (n = 202)</th>
<th>Control group (n = 204)</th>
<th>Incremental costs</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Intervention</strong></td>
<td>Mean, € (SD)</td>
<td>Mean, € (SD)</td>
<td>Difference, €</td>
</tr>
<tr>
<td></td>
<td>299 (-)</td>
<td>10 (-)</td>
<td>289</td>
</tr>
<tr>
<td><strong>Health care costs</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>GP/internist</td>
<td>142 (142)</td>
<td>117 (154)</td>
<td>25</td>
</tr>
<tr>
<td>Mental health care</td>
<td>117 (308)</td>
<td>175 (447)</td>
<td>-58</td>
</tr>
<tr>
<td>Other medical specialist a</td>
<td>243 (428)</td>
<td>236 (363)</td>
<td>7</td>
</tr>
<tr>
<td>In-patient care</td>
<td>48 (679)</td>
<td>123 (1302)</td>
<td>-75</td>
</tr>
<tr>
<td>Day care</td>
<td>0 (-)</td>
<td>35 (499)</td>
<td>-35</td>
</tr>
<tr>
<td>Antidepressants</td>
<td>12 (41)</td>
<td>20 (57)</td>
<td>-8</td>
</tr>
<tr>
<td><strong>Patient and family costs</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Private therapist b</td>
<td>137 (354)</td>
<td>117 (246)</td>
<td>20</td>
</tr>
<tr>
<td>Co-payments c</td>
<td>27 (73)</td>
<td>30 (65)</td>
<td>-3</td>
</tr>
<tr>
<td>Over-the-counter drugs</td>
<td>17 (34)</td>
<td>22 (42)</td>
<td>-5</td>
</tr>
<tr>
<td>Informal care</td>
<td>323 (943)</td>
<td>384 (857)</td>
<td>-61</td>
</tr>
<tr>
<td>Domestic help</td>
<td>143 (455)</td>
<td>120 (511)</td>
<td>-23</td>
</tr>
<tr>
<td>Travel</td>
<td>29 (77)</td>
<td>28 (72)</td>
<td>1</td>
</tr>
<tr>
<td><strong>Productivity losses</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Absenteeism</td>
<td>1475 (2498)</td>
<td>1172 (2209)</td>
<td>303</td>
</tr>
<tr>
<td>Presenteeism</td>
<td>1696 (1622)</td>
<td>2021 (2781)</td>
<td>-325</td>
</tr>
<tr>
<td><strong>Total health care costs</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Total</td>
<td>904 (989)</td>
<td>768 (1777)</td>
<td>136</td>
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<tr>
<td></td>
<td>4655 (4674)</td>
<td>4513 (5160)</td>
<td>143</td>
</tr>
</tbody>
</table>

Abbreviations: €, costs in Euro; SD, standard deviation

a i.e., physiotherapist, occupational therapist
b i.e., physiotherapist without prescription
c i.e., patient’s contribution to prescribed medication
Societal perspective

Table 2 shows the incremental cost, effects and cost-effectiveness ratios (based on 2,500 bootstrap simulations) for the main analysis and the sensitivity analyses, and from both the societal and health systems perspective. From a societal perspective, the intervention resulted in a greater mean health benefit (0.12 DFY gained) achieved at higher mean total costs (€134; £114) as compared to enhanced CAU.

The cost-effectiveness plane, representing the 2,500 bootstrap replications, is shown in Figure 1. Most (62%) of the bootstrapped ICERs fell in the North East quadrant, indicating a 62% probability that the intervention produces greater health, but at greater costs than enhanced CAU. The remaining 38% of ICERs fell in the South East quadrant, indicating a 38% probability that the intervention dominates enhanced CAU because additional health gains are obtained for lesser costs. In other words, at a willingness-to-pay (WTP) of €0, the probability that the intervention must be regarded as more cost-effective than CAU is 38%. However, at a WTP of €7,350 (£6,248), €9,680 (£8,228) and €20,000 (£17,000) for gaining a depression-free life-year, the intervention’s probability of being more cost-effective than CAU rises to 90%, 95%, and 99% (Figure 2). The ICER based on QALY gains showed a small health benefit (0.01 QALYs gained) for higher mean costs (€134; £114). As seen in Figure 3, most of the simulated ICERs fell in the North East quadrant (49%) (see also Table 2). The intervention’s probability of dominating enhanced CAU was 35% when taking the societal perspective. Assuming a willingness-to-pay of €20,000 (£17,000) for gaining a QALY, this probability rose to 60% (Figure 4).
### Table 2. Results of the main and sensitivity analyses (based on 2,500 bootstrap simulations)

<table>
<thead>
<tr>
<th>Analysis and perspective</th>
<th>Incremental costs, € (95% CI)</th>
<th>Incremental effects (95% CI)</th>
<th>Mean ICER (95% CI)</th>
<th>Distribution over the ICER plane</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
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<td>NE</td>
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<tr>
<td><strong>Cost-effectiveness, DFYs</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Societal</td>
<td>134 (-827 - 1,055)</td>
<td>0.12 (0.05 - 0.18)</td>
<td>1,117 (-7,546 - 11,737)</td>
<td>62%</td>
</tr>
<tr>
<td>Health care</td>
<td>135 (-146 - 418)</td>
<td>0.12 (0.05 - 0.18)</td>
<td>1,125 (-1,428 - 4,715)</td>
<td>83%</td>
</tr>
<tr>
<td><strong>Cost-utility, EQ-5D QALYs</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Societal</td>
<td>134 (-827 - 1,055)</td>
<td>0.01 (-0.01 - 0.04)</td>
<td>13,400 (^a)</td>
<td>49%</td>
</tr>
<tr>
<td>Health care</td>
<td>135 (-146 - 418)</td>
<td>0.01 (-0.01 - 0.04)</td>
<td>13,500 (^a)</td>
<td>68%</td>
</tr>
<tr>
<td><strong>Sensitivity analyses</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>SF-6D QALYs, societal</td>
<td>134 (-827 - 1,055)</td>
<td>0.03 (0.02 - 0.05)</td>
<td>4,467 (-23,846 - 42,891)</td>
<td>60%</td>
</tr>
<tr>
<td>SF-6D QALYs, health care</td>
<td>135 (-146 - 418)</td>
<td>0.03 (0.02 - 0.05)</td>
<td>4,500 (-5000 - 15,088)</td>
<td>83%</td>
</tr>
<tr>
<td><strong>Without hospitalization costs</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Societal, DFY</td>
<td>233 (-649 - 1,155)</td>
<td>0.12 (0.05 - 0.18)</td>
<td>1,942 (-5169 - 14,705)</td>
<td>70%</td>
</tr>
<tr>
<td>Health care, DFY</td>
<td>245 (119 - 374)</td>
<td>0.12 (0.05 - 0.18)</td>
<td>2,042 (865 - 5,562)</td>
<td>100%</td>
</tr>
<tr>
<td>Societal, EQ-5D QALY</td>
<td>233 (-649 - 1,155)</td>
<td>0.01 (-0.01 - 0.04)</td>
<td>23,300 (^a)</td>
<td>13%</td>
</tr>
<tr>
<td>Health care, EQ-5D QALY</td>
<td>245 (119 - 374)</td>
<td>0.01 (-0.01 - 0.04)</td>
<td>24,500 (4,220 - 22,333)</td>
<td>16%</td>
</tr>
</tbody>
</table>

Abbreviations: 95% CI, 95% confidence interval; DFY, Depression-free years; EQ-5D QALYs, quality-adjusted life years based on EuroQol; SF-6D QALYs, quality-adjusted life years based on SF-12; ICER, incremental cost-effectiveness ratio; NE, northeast quadrant; NW, northwest quadrant; SE, southeast quadrant; SW, southwest quadrant.

\(^a\) A dependably accurate 95% confidence interval for this distribution cannot be defined because there is no line through the origin that excludes \(\alpha/2\) of the distribution.
Figure 1. Scatterplot of 2,500 replicates of the incremental cost-effectiveness ratio (mean differences in costs from a societal perspective and in depression-free years) on the cost-effectiveness plane: web-based guided self-help intervention vs. enhanced usual care.

Figure 2. Cost-effectiveness acceptability curve showing the probability of the web-based guided self-help intervention being cost-effective at varying willingness-to-pay ceilings (based on 2,500 replicates of the incremental cost-effectiveness ratio using mean differences in costs from a societal perspective and depression-free years).
**Figure 3.** Scatterplot of 2,500 replicates of the incremental cost-effectiveness ratio (mean differences in costs from a societal perspective and in quality-adjusted life years (QALYs)) on the cost-effectiveness plane: web-based guided self-help intervention vs. enhanced usual care

**Figure 4.** Cost-effectiveness acceptability curve showing the probability of the web-based guided self-help intervention being cost-effective at varying willingness-to-pay ceilings (based on 2,500 replicates of the incremental cost-effectiveness ratio using mean differences in costs from a societal perspective and quality-adjusted life years (QALYs))
Public health care perspective

From a public health care perspective, the health benefit (0.12 DFY gained) was also achieved at higher costs (£135; £115). In the cost-effectiveness plane, 83% of simulated ICERs fell in the North East quadrant with the remaining 17% in the South East quadrant (Table 2). The intervention’s probability to be more cost-effective than enhanced CAU increased to 51% at a WTP of £1,250 (£1063) and became 90%, 95%, and 99% at a WTP of £3,140 (£2669), £3,920 (£3332), and £6,000 (£5100), respectively.

Like from the societal perspective, the ICER from the public health care perspective based on QALY gains showed a small health benefit (0.01 QALYs gained) for higher mean costs (£135). In the cost-effectiveness plane, 68% of the simulated ICERs fell in the North East quadrant (Table 2). The intervention’s probability of dominating enhanced CAU was 16% when considering health care costs. Assuming a willingness-to-pay of £20,000 (£17,000) for gaining a QALY, this probability was 64%.

Sensitivity analyses

Using the SF-6D resulted in a greater incremental QALY gain in favour of the intervention group (0.70 QALY, SD = 0.08) as compared to the control group (0.67 QALY, SD = 0.07), which was statistically significant [t(404) = -4.40, p < .001]. This agrees with available evidence that the EQ-5D suffers from a ceiling effect in less severe diseases (43). At a WTP of £20,000 for gaining one QALY the probability of being cost-effective was 84% (societal perspective) and 99% (health care perspective). Hospital costs were higher in the control group so excluding these costs resulted in higher ICERs (Table 2). From a societal perspective and at a WTP of £20,000, the intervention’s probability of being cost-effective then became 53%.

Discussion

Main findings

Our study was set out to evaluate the cost-effectiveness and cost-utility of a web-based guided self-help intervention to prevent the onset of major depressive disorder (MDD) in adults suffering from subthreshold depression in comparison to usual care enhanced with web-based psycho-education. Main outcomes were depression-free life years (DFYs) and quality adjusted life years (QALYs).

Significantly more depression-free years (DFYs) were gained in the intervention group as compared to the control group. The probability that the intervention is deemed to be cost-effective depends on the willingness-to-pay (WTP) for a depression-free year. Assuming that the
intervention is considered to be cost-effective if its likelihood of being cost-effective is greater than 90%, a WTP of €7,350 and above for gaining a DFY would make the intervention cost-effective. While these DFY health gains were not mirrored in EQ-5D QALY gains, they were observed in SF-6D QALY gains with the latter being more sensitive to change in milder conditions such as subthreshold depression and therefore not suffering as much from a ceiling effect as the EQ-5D does (31). If society would be willing to pay €20,000 (£17000) for gaining a QALY, the probability of being cost-effective will be 64% for gaining an EQ-5D QALY and 84% for gaining a SF-6D QALY.

Wider context
Both burden of disease studies and cost of illness studies (1, 2) attest to the importance of cost-effective interventions that can reduce the burden of depressive disorders on a wider scale. Recent meta-analytic evidence suggests that it is possible to prevent the onset of major depression using psychological interventions (8). Results of our study showed that a guided E-health intervention can successfully reduce the incidence of diagnosed major depression (16). In addition, some evidence indicates that bibliotherapeutic self-help interventions for the prevention of depression represent good value-for-money (15).

However, economic evaluations in the field of depression prevention mainly relied on health-economic modelling (12, 13, 44) with direct evidence stemming only from two randomised controlled trials (10, 11). Although some trial-based economic evaluations of computerised cognitive-behavioural interventions for treating depressive symptoms exist (i.e., (45, 46)), to our knowledge, this is the first trial-based economic evaluation of a web-based intervention to prevent the onset of major depression in an adult population with subthreshold depression. Results from our trial adds to the converging evidence pointing to the effectiveness and cost-effectiveness (sometimes even costs-savings) of depression prevention across a range of outcomes.

Limitations
This study has some limitations. Firstly, the time horizon of this study was limited to 12 months. A recent meta-analysis showed a small positive association between effects of preventive interventions during the first months of follow-up, indicating that with passing months, the intervention effects get somewhat larger. However, meta-regression analyses suggested that the effects of interventions are lower at longer follow-up periods of one to two years (8). Assuming diminishing long-term effects, the cost-effectiveness of this particular intervention will also decline. However, only few studies had longer follow-up periods than 2 years. Thus,
more research with longer follow-up periods are needed to assess the long-term effectiveness and cost-effectiveness of preventive (web-based) interventions.

Secondly, we did not assess lifetime history of MDD at baseline and therefore we cannot be sure if we prevented first-ever onsets of MDD or MDD recurrences. Future studies should thus clarify whether web-based guided self-help interventions are cost-effective both for the prevention of first depression onset and the prevention of recurrence.

Thirdly, costs were assessed with the help of self-reports and may suffer from under-reporting. However, the structured questionnaire used in the current study can be considered as valid instrument for recall periods up to 3 months (47).

Finally, the trial has been conducted in a highly educated sample. Evidence suggests that better adherence is predicted by higher education (48). In our trial, only 2% of participants were low educated. Hence, we cannot predict the uptake of such an intervention in less educated people or among people with a lower socioeconomic status. However, we used an open recruitment strategy in our trial mimicking the way how people will be recruited for e-health interventions in the future, thus providing ecological validity to the sample on which the current study is based. In other words, the trial sample reflects the population segments that are interested in engaging in a web-based intervention. However, one conclusion drawn from this trial is that not all people who are in need of psychological interventions could be reached via the Internet. The applicability of web-based interventions is related to (a) the acceptance of such interventions by the target population (i.e. preferences for different treatment modalities, such as face-to-face interventions) and (b) the availability of technical requirements (i.e. reliable access to the Internet).

Clinical implications and future research
Current guidelines on depression treatment (such as the NICE guideline and the Dutch Multidisciplinary Guideline for Depressive Disorder) recommend low-intensity psychosocial interventions (i.e., computerised cognitive behavioural therapy) to manage (persistent) subthreshold depressive symptoms and mild to moderate depression (22). Our study supports this recommendation by showing that an E-health intervention may not only restore health in people with subthreshold depression, but in addition reduces the risk of developing a major depressive disorder. Findings from our study also adds that delivering cognitive-behavioural therapy over the Internet has a high probability of being cost-effective to prevent the onset of new depressive episodes. Given low participation rates in face-to-face preventive services and the potential to scale up web-based interventions to efficiently alleviate the disease burden caused by MDD, it would be worthwhile to integrate such a web-based intervention into routine practice.
However, there are some opportunities and risks that need to be taken into account when scaling up this intervention. Firstly, the individual online-trainer feedback provided in this web-based intervention hinders scaling up the intervention. A recent review of randomised controlled trials showed that unguided interventions can also be effective (with lower adherence rates compared to guided interventions (49)). To be more precise, in our study, out of the 202 participants who were initially assigned to the intervention, 138 (68.3%) were intervention completers. This compares favourably to an unguided Internet-based intervention for the treatment of sub-threshold depression that was completed by 48.3% (49/102) of participants (50). Providing guidance may not only affect the outcome and cost-effectiveness of the intervention but also the target group’s willingness to use such an intervention and thereby influencing the effects of such an intervention at population level. It is therefore not possible to predict the effects of this particular web-based intervention at population level when it is offered without guidance by online-trainers. In addition, there are no guarantees that adherence and (by proxy) effectiveness will be maintained if web-based preventive interventions are scaled up in the population. For example, Christensen and colleagues reported that less than 1% of public registrants using a preventive intervention delivered openly on a website completed all modules (51).

Secondly, an unanswered question refers to how web-based interventions could be rolled out to the population. For example, in the UK and the Netherlands, Internet-delivered CBT is already prescribed by GPs (22). In addition, web-based interventions could be promoted independently at the GP’s (i.e., promotion videos in the GP’s waiting rooms). However, one challenge in depression prevention is that most individuals at risk of developing a major depression do not show up in primary care. Therefore, innovative approaches are needed to reach these groups, for example through a systematic mental health screening of all people in specific settings (i.e., occupational setting or universities) and to motivate those at risk to engage in preventive interventions (i.e., acceptance facilitating interventions (52)). However, such strategies do not guarantee uptake either.

Thirdly, the cost of the web-based intervention were calculated without considering an economies of scale effect. Economies of scale refer to the reduction in the cost per treatment of an intervention as a result of increasing the number of clients. Economies of scale arise because many of the costs associated with the web-based intervention are fixed and not dependent on the number of clients (i.e., hosting the intervention on a server) and thus increasing the intervention output reduces the fixed cost per treatment. Hence, we assume that the cost of IT infrastructure (i.e., the per-client cost of servers to host the intervention) will be cheaper per additional client if the intervention is offered on a larger scale. Economies of scale might also reduce variable costs (i.e., therapist’s support per participant) because therapists
become more efficient through better organisation and learning by doing. However, the same technical resources available in the research setting (i.e., reasonable Internet connections) may not be available when the intervention is scaled up.

Finally, for some individuals a self-help approach might not be sufficient (53). Some individuals may feel unable to apply psychotherapeutic self-help strategies. Some techniques could be inappropriately implemented by participants without guidance by an online-trainer. One could argue that it is easier to observe and react to early signs of deterioration in face-to-face interventions as compared to web-based interventions. Another potential negative effect of self-help interventions could be a delay in help-seeking leading to a further deterioration of symptoms, if the initial low-intensity self-help intervention should not be sufficient. Hence, multiple approaches to reach the target population are needed in successful depression prevention programmes. Future studies should thus evaluate the preventive effects of unguided web-based interventions on the onset of major depressive disorder and compare the cost-effectiveness of unguided and guided interventions. In addition, implementation studies should be conducted to obtain real world effects of such interventions and to gather knowledge about the willingness to use these interventions in specific population segments (i.e., in low educated people, in a rural setting or among people with lower socioeconomic status).

**Conclusion**

Given the evidence for the efficacy of psychological interventions to prevent depression and the potential scalability and cost-effectiveness of web-based interventions, large-scale dissemination of these interventions might be a promising strategy to alleviate depression’s disease burden in an affordable way and on a wide scale.

However, before a nationwide dissemination could be considered, future studies need to evaluate the preventive effects of unguided web-based interventions on the onset of major depressive disorder and compare the cost-effectiveness of unguided and guided interventions. Moreover, implementation studies are needed to obtain real world effects of such interventions and to gather knowledge about the willingness to use these interventions in specific population segments (i.e., in low educated people, in a rural setting or among people with lower socioeconomic status).
References


